

# 7. Cross tabulation and logistic regression

## (PROC FREQ and PROC LOGISTIC)

Use of SAS  
March 2011

## Table analyses: PROC FREQ

```
/* One-way table */  
proc freq data=afrika.bissau;  
    tables dead;  
run;
```

```
/* Two-way table */  
proc freq data=afrika.bissau;  
    tables bcg*dead;  
run;
```

```
/* Three-way table */  
proc freq data=afrika.bissau;  
    tables agemm*bcg*dead;  
run;
```

*etc.*

# Two-way table (2 x 2 table)

```
proc freq data=afrika.bissau;
  tables bcg*dead;
run;
```

Table of bcg by dead

bcg	dead		
Frequency			
Percent			
Row Pct			
Col Pct	1	2	Total
-----+-----+-----+			
1	124	3176	3300
	2.35	60.23	62.58
	3.76	96.24	
	56.11	62.87	
-----+-----+-----+			
2	97	1876	1973
	1.84	35.58	37.42
	4.92	95.08	
	43.89	37.13	
-----+-----+-----+			
Total	221	5052	5273
	4.19	95.81	100.00

## Two-way table (2 x 2 table)

Exposure	Outcome		Total
	Yes	No	
Yes	$a$	$b$	$n_1$
No	$c$	$d$	$n_2$
Total	$a + c$	$b + d$	$n$

Hypothesis  $H_0$ : The probability of having the outcome is the same in the two exposure groups.

The probability of having the outcome under  $H_0$ :

$$p = \frac{a + c}{n}$$

.

The EXPECTED numbers under  $H_0$  in the four cells are calculated as:

Exposure	Outcome		Total
	Yes	No	
Yes	$E(a) = p \times n_1$	$E(b) = (1 - p) \times n_1$	$n_1$
No	$E(c) = p \times n_2$	$E(d) = (1 - p) \times n_2$	$n_2$
Total	$a + c$	$b + d$	$n$

Chi-square test for testing  $H_0$  (observed - expected):

$$X^2 = \frac{(a - E(a))^2}{E(a)} + \frac{(b - E(b))^2}{E(b)} + \frac{(c - E(c))^2}{E(c)} + \frac{(d - E(d))^2}{E(d)}$$

$$\sim \chi^2(1)$$

$H_0$  is rejected if p-value  $< 0.05$  which corresponds to  $X^2 > 3.84$ .

## Risk of Dying and BCG

```
proc freq data=afrika.bissau;  
  tables bcg*dead / expected chisq nocol nopercent;  
run;
```

Table of bcg by dead

bcg	dead		
Frequency			
Expected			
Row Pct	1	2	Total
-----+-----+-----+			
1	124	3176	3300
	138.31	3161.7	
	3.76	96.24	
-----+-----+-----+			
2	97	1876	1973
	82.692	1890.3	
	4.92	95.08	
-----+-----+-----+			
Total	221	5052	5273

Statistics for Table of bcg by dead

Statistic	DF	Value	Prob
-----			
Chi-Square	1	4.1291	0.0422
Likelihood Ratio Chi-Square	1	4.0516	0.0441
Continuity Adj. Chi-Square	1	3.8456	0.0499
Mantel-Haenszel Chi-Square	1	4.1283	0.0422
Phi Coefficient		-0.0280	
Contingency Coefficient		0.0280	
Cramer's V		-0.0280	

## Risk Ratio

Exposure	Outcome		Total
	Yes	No	
Yes	$a$	$b$	$n_1$
No	$c$	$d$	$n_2$
Total	$a + c$	$b + d$	$n$

Risk ratio:

$$\text{RR} = \frac{\text{probability of outcome among exposed}}{\text{probability of outcome among not-exposed}} = \frac{a/n_1}{c/n_2}.$$

The  $H_0$  corresponds to  $\text{RR} = 1$ .



## Odds ratio

Exposure	Outcome		Total
	Yes	No	
Yes	$a$	$b$	$n_1$
No	$c$	$d$	$n_2$
Total	$a + c$	$b + d$	$n$

Odds ratio:

$$\text{OR} = \frac{\text{odds of outcome among exposed}}{\text{odds of outcome among not-exposed}} = \frac{a/b}{c/d} = \frac{a \times d}{b \times c}$$

The  $H_0$  corresponds to  $\text{OR} = 1$ .

# RR and OR in proc freq

```
proc freq data=afrika.bissau;
  table bcg*dead / RELRISK nocol nopercent;
run;
```

Table of bcg by dead

bcg		dead		
Frequency				
Row	Pct	1	2	Total
-----+-----+-----+				
1		124	3176	3300
		3.76	96.24	
-----+-----+-----+				
2		97	1876	1973
		4.92	95.08	
-----+-----+-----+				
Total		221	5052	5273

Statistics for Table of bcg by dead

Estimates of the Relative Risk (Row1/Row2)

Type of Study	Value	95% Confidence Limits	
-----			
Case-Control (Odds Ratio)	0.7551	0.5754	0.9909
Cohort (Col1 Risk)	0.7643	0.5895	0.9910
Cohort (Col2 Risk)	1.0122	1.0000	1.0245

## OR and RR in proc freq

It is important that the two variables in the `table` statement are coded properly when using the OR and RR:

```
data hope;
  set afrika.bissau;
  if dead=2 then deadny=0;
  if dead=1 then deadny=1;
proc freq data=hope;
  table bcg*deadny / relrisk nocol nopercent norow;
run;
```

bcg		deadny		
Frequency				
Row	Pct	0	1	Total
-----+-----+-----+				
1		3176	124	3300
-----+-----+-----+				
2		1876	97	1973
-----+-----+-----+				
Total		5052	221	5273

Statistics for Table of bcg by deadny			
Estimates of the Relative Risk (Row1/Row2)			
Type of Study	Value	95% Confidence Limits	
-----			
Case-Control (Odds Ratio)	1.3243	1.0092	1.7378
Cohort (Col1 Risk)	1.0122	1.0000	1.0245
Cohort (Col2 Risk)	0.7643	0.5895	0.9910

## R x C tables

```
proc freq data=afrika.bissau;  
  table ethnic*dead / chisq;  
run;
```

ethnic	dead		
Frequency			
Percent			
Row Pct			
Col Pct	1	2	Total
-----+-----+-----+			
Balanta	37	788	825
	0.70	14.94	15.65
	4.48	95.52	
	16.74	15.60	
-----+-----+-----+			
Fula	52	1370	1422
	0.99	25.98	26.97
	3.66	96.34	
	23.53	27.12	
-----+-----+-----+			
Mandinga	49	1113	1162
	0.93	21.11	22.04
	4.22	95.78	
	22.17	22.03	
-----+-----+-----+			
Other	23	724	747
	0.44	13.73	14.17
	3.08	96.92	
	10.41	14.33	
-----+-----+-----+			
Pepel	60	1057	1117
	1.14	20.05	21.18
	5.37	94.63	
	27.15	20.92	
-----+-----+-----+			
Total	221	5052	5273
	4.19	95.81	100.00

Statistics for Table of ethnic by dead

Statistic	DF	Value	Prob
Chi-Square	4	7.3670	0.1177
Likelihood Ratio Chi-Square	4	7.3268	0.1196
Mantel-Haenszel Chi-Square	1	1.0857	0.2974
Phi Coefficient		0.0374	
Contingency Coefficient		0.0374	
Cramer's V		0.0374	

## Exercise: PROC FREQ

Using the bissau data:

1. Investigate whether DTP-vaccinated children (variable `ntp`) dies more often than DTP-unvaccinated children.
2. Calculate the odds ratio (OR) and corresponding 95% confidence interval.
3. The variable `region` indicates the rural region of the children. Is mortality associated with region?

## Logistic regression: PROC LOGISTIC

Logistic regression is like a linear regression, but here the outcome is DISCRETE with two levels (yes/no, died/survived, ill/well).

Look again at the 2 x 2 table

Exposure	Outcome		Total
	Yes	No	
Yes	$a$	$b$	$n_1$
No	$c$	$d$	$n_2$

Let  $p = a/n_1$  be the probability of outcome among exposed. Odds can then be defined as

$$\text{odds} = \frac{p}{1-p} = \frac{a/n_1}{1-a/n_1} = \frac{a/n_1}{b/n_1} = \frac{a}{b}$$



## Logistic regression for 2 x 2 table

What is modeled in a logistic regression is the NATURAL LOGARITHM of the ODDS of outcome:

$$\ln(\text{odds}) = \ln \left( \frac{p}{1-p} \right) = \beta_0 + \beta_1 X,$$

where  $X$  is the exposure covariate. We call  $\ln(\text{odds})$  for the LOG-ODDS. Now let us assume that the exposure is coded like

$$X = \begin{cases} 1 & \text{Exposed} \\ 0 & \text{Non-exposed} \end{cases}$$

The log-odds of outcome among exposed ( $X = 1$ ) is

$$\ln \left( \frac{p_1}{1 - p_1} \right) = \beta_0 + \beta_1 \times 1 = \beta_0 + \beta_1.$$

The log-odds of outcome among non-exposed ( $X = 0$ ) is

$$\ln \left( \frac{p_0}{1 - p_0} \right) = \beta_0 + \beta_1 \times 0 = \beta_0.$$

The  $\beta_0$  is the log-odds of outcome in non-exposed.

Now, the difference in log-odds between exposed and non-exposed is

$$\ln \left( \frac{p_1}{1 - p_1} \right) - \ln \left( \frac{p_0}{1 - p_0} \right) = \beta_0 + \beta_1 - \beta_0 = \beta_1$$

Using the rule of logarithms

$$\ln(a) - \ln(b) = \ln\left(\frac{a}{b}\right)$$

we get

$$\ln\left(\frac{p_1/(1-p_1)}{p_0/(1-p_0)}\right) = \beta_1$$

and

$$\frac{p_1/(1-p_1)}{p_0/(1-p_0)} = \exp(\beta_1)$$

This means that the odds ratio between exposed and non-exposed is

$$\text{OR} = \exp(\beta_1).$$

Estimation of the regression coefficients is done using maximum likelihood.

# PROC LOGISTIC

```
proc logistic data=afrika.bissau;  
  class bcg / param=ref;  
  model dead(event="1")=bcg;  
run;
```

**REMEMBER** the option param=ref

Part of output:

Analysis of Maximum Likelihood Estimates						
Parameter	DF	Estimate	Standard Error	Chi-Square	Pr > ChiSq	
Intercept	1	-2.9621	0.1041	809.3011	<.0001	
bcg	1	-0.2810	0.1386	4.1074	0.0427	

Odds Ratio Estimates			
Effect	Point Estimate	95% Wald Confidence Limits	
bcg 1 vs 2	0.755	0.575	0.991

## Multiple logistic regression

$$\ln(\text{odds}) = \ln \left( \frac{p}{1-p} \right) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \cdots ,$$

The interpretation is still that  $\exp(\beta_1)$  is an odds ratios, but now adjusted for the covariates  $X_2, X_3, \cdots$ . The same idea as in linear regression.

Remember that the response or outcome is discrete with two categories. However the covariates  $(X_1, X_2, X_3, \cdots)$  do not need to be categorical, they can also be continuous.

In SAS one use the CLASS statement to indicate categorical variables, and variables in a MODEL statement that is not listed in the CLASS statement is assumed to be continuous.

# Multiple logistic regression: PROC LOGISTIC

```
proc logistic data=afrika.bissau;
  class bcg / param=ref;
  model dead(event="1")=bcg agemm;
run;
```

## Type 3 Analysis of Effects

Effect	DF	Wald	
		Chi-Square	Pr > ChiSq
bcg	1	5.4366	0.0197
agemm	1	1.5307	0.2160

## Analysis of Maximum Likelihood Estimates

Parameter	DF	Estimate	Standard		Wald	
			Error		Chi-Square	Pr > ChiSq
Intercept	1	-3.0510	0.1281		567.4989	<.0001
bcg 1	1	-0.3450	0.1480		5.4366	0.0197
agemm	1	0.0486	0.0393		1.5307	0.2160

## Odds Ratio Estimates

Effect		Point		95% Wald	
		Estimate		Confidence Limits	
bcg	1 vs 2	0.708		0.530	0.946
agemm		1.050		0.972	1.134

Interpretation: For each increase of 1 in **agemm** the odds increases with 1.050.

## Multiple logistic regression: PROC LOGISTIC

The variable `agemm` is now used as a CLASS variable:

```
proc logistic data=afrika.bissau;  
  class bcg agemm / param=ref;  
  model dead(event="1")=bcg agemm;  
run;
```

`agemm` has 7 classes: 0 to 6. SAS automatically generates 7 indicator functions for each class and include 6 of these in the regression model. The class not included (SAS uses per default the highest class) is the reference against which all other classes are listed in the output.

The test in TYPE 3 for `agemm` is a test for the hypothesis of equal risk of dying in the 7 classes. This test does not change when you change reference.

### Type 3 Analysis of Effects

Wald			
Effect	DF	Chi-Square	Pr > ChiSq
bcg	1	5.2393	0.0221
agemm	6	7.3938	0.2860

### Analysis of Maximum Likelihood Estimates

		Standard		Wald	
Parameter	DF	Estimate	Error	Chi-Square	Pr > ChiSq
Intercept	1	-3.0948	0.3281	88.9498	<.0001
bcg	1	-0.3430	0.1499	5.2393	0.0221
agemm	0	0.0588	0.3599	0.0267	0.8703
agemm	1	0.1803	0.3513	0.2635	0.6077
agemm	2	-0.1925	0.3650	0.2783	0.5978
agemm	3	0.2700	0.3514	0.5904	0.4423
agemm	4	0.4044	0.3492	1.3410	0.2469
agemm	5	0.3618	0.3549	1.0392	0.3080

### Odds Ratio Estimates

		Point	95% Wald	
Effect		Estimate	Confidence Limits	
bcg	1 vs 2	0.710	0.529	0.952
agemm	0 vs 6	1.061	0.524	2.147
agemm	1 vs 6	1.198	0.602	2.384
agemm	2 vs 6	0.825	0.403	1.687
agemm	3 vs 6	1.310	0.658	2.608
agemm	4 vs 6	1.498	0.756	2.971
agemm	5 vs 6	1.436	0.716	2.879



## Change of reference group: REF=""

The variable `agemm` is again used as a CLASS variable but now choosing agegroup 4 as reference:

```
proc logistic data=afrika.bissau;  
  class bcg agemm(ref="4") / param=ref;  
  model dead(event="1")=bcg agemm;  
run;
```

### Type 3 Analysis of Effects

Effect	DF	Wald	
		Chi-Square	Pr > ChiSq
bcg	1	5.2393	0.0221
agemm	6	7.3938	0.2860

### Analysis of Maximum Likelihood Estimates

Parameter	DF	Estimate	Standard Error	Wald	
				Chi-Square	Pr > ChiSq
Intercept	1	-2.6904	0.1973	185.9085	<.0001
bcg	1	-0.3430	0.1499	5.2393	0.0221
agemm	0	-0.3456	0.2479	1.9444	0.1632
agemm	1	-0.2241	0.2368	0.8953	0.3441
agemm	2	-0.5969	0.2575	5.3731	0.0204
agemm	3	-0.1344	0.2385	0.3175	0.5731
agemm	5	-0.0426	0.2442	0.0304	0.8616
agemm	6	-0.4044	0.3492	1.3410	0.2469

### Odds Ratio Estimates

Effect		95% Wald	
		Point Estimate	Confidence Limits
bcg	1 vs 2	0.710	0.529 0.952
agemm	0 vs 4	0.708	0.435 1.150
agemm	1 vs 4	0.799	0.502 1.271
agemm	2 vs 4	0.551	0.332 0.912
agemm	3 vs 4	0.874	0.548 1.395
agemm	5 vs 4	0.958	0.594 1.546
agemm	6 vs 4	0.667	0.337 1.323

## Exercise: PROC LOGISTIC

Using the Bissau data:

1. Make a logistic regression where outcome is `dead` and exposure is `ntp`. Interpret the results and compare with the results from the exercise using `proc freq` on page 14.
2. Now control for `bcb` in the logistic regression from 1 above. What happen with the odds ratio for `ntp`?
3. Add variables `agemm` and `region` to the model as class variables. Let `region=7` be the reference group for variable `region`. Did inclusion of these variables change interpretation of effect `ntp`?

## Logistic regression: PROC GENMOD

The procedure `proc genmod` (GENeralized linear MODels) can also perform logistic regression:

```
proc genmod data=afrika.bissau;  
  class bcg;  
  model dead=bcg / dist=binomial;  
  estimate "BCG+ vs BCG-" bcg 1 -1 / exp;  
run;
```

Getting the odds ratios and their confidence intervals out is however not as easy as in `proc logistic`, but done with `estimate` statements. `proc genmod` is however very flexible and can analyse other types of regression models like Poisson (event per person-years) and linear regression.